Introduction about Copulas.

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The talk is based on two following references:

- An Introduction to Copulas. (2006) Roger B. Nelsen, Springer.
- Elements of Copula Modeling with R. (2018) M. Hofert, I.Kojadinovic, M. Machler, J. Yan, Springer.

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Outline

- First part: Motivating Example.
- Second part: Analysis Approach to Copula.
 - Subcopula and Copula
 - Sklar's Theorem.
- Third part: Copulas and Random variables.
 - Sklar's Theorem.
 - Fréchet-Hoeffding Bounds
- Forth part: Survival Copula and Symmetric Copula.

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PART 1: Motivating Example.

The data set.

Let $(X_1, X_2)^{(n)} = (X_{i_1}, X_{i_2})_{i=1}^n$, $(Y_1, Y_2)^{(n)} = (Y_{i_1}, Y_{i_2})_{i=1}^n$ be two given data set.

Goal: Comparing the dependence between between X_1 and X_2 with Y_1 and Y_2 .



Figure: n = 1000 independent observations of (X_1, X_2) and (Y_1, Y_2) .

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Kernel Estimation.



Figure: Kernel density estimates of the densities of (X_1, X_2) and (Y_1, Y_2) .

• X₁ and X₂ are likely to follow standard normal distribution.

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• Y_1 and Y_2 are likely to follow standard exponential distribution.

Transform to the uniform distributions.

Lemma: Let X be continuous random variable with distribution function F, then F(X) is a standard uniform random variable, i.e., $F(X) \sim U(0,1)$



Figure: Scatter plots of $(F(X_1), F(X_2))$ and $(G(Y_1), G(Y_2))$

▶ The distribution of $(F(X_1), F(X_2))$ and $(G(Y_1), G(Y_2))$ seem to be identical.

► Two given data sets is indistinguishable in terms of **dependence** and only differ in terms of the underlying **marginal distribution function**.

Question: Is there any tool to measure the **dependence** between random variables?

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PART 2: Analysis approach to Copula.

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• A rectangle
$$K$$
 in $\overline{\mathbb{R}}^2$ is denoted by:

$$\boldsymbol{\mathsf{K}}=[\boldsymbol{x}_1,\boldsymbol{x}_2]\times[\boldsymbol{y}_1,\boldsymbol{y}_2].$$

For any function $H : \mathbb{R}^2 \to \mathbb{R}$, **DomH** is its domain and **RanH** is its range. Assume that **DomH** = $S_1 \times S_2$.

• Denote a_1, a_2 be the least elements of S_1, S_2 respectively and b_1, b_2 be the greastest elements of S_1, S_2 respectively

► *H* – *volume* of *K*:

 $V_{H}(K) = H(x_{2}, y_{2}) - H(x_{2}, y_{1}) - H(x_{1}, y_{2}) + H(x_{1}, y_{1})$

• **H** deduces a measure on $\overline{\mathbb{R}}^2$.

▶ A 2 - place real-valued function H is 2 - increasing function if $V_H(K) \ge 0$ for all rectangles within *DomH*.

▶ Let a_1, a_2 be the least elements of S_1, S_2 respectively. The function $H: S_1 \times S_2 \rightarrow \mathbb{R}$ is grounded if

 $H(x,a_2)=0=H(a_1,y)$ for all $(x,y)\in S_1\times S_2$.

Remark:

2 - increasing \Rightarrow grounded : H(x, y) = (2x - 1)(2y - 1)grounded \Rightarrow 2 - increasing : $H(x, y) = \max(x, y)$

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Lemma: If H be a grounded, 2 - increasing function H is nondecreasing in each argument.

► A two - dimensional real-valued subcopula *C*' is a function which has the following properties:

• DomC' is an subset of square unit \mathbb{I}^2 containing the points 0 and 1.

- C' is grounded and 2 increasing.
- For every (u, v) in $S_1 \times S_2$, C'(u, 1) = u, C'(1, v) = v.

▶ A two - dimensional real-valued **copula** is a **subcopula** whose domain is \mathbb{I}^2 .

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► Considering the **Frank** Copula

$$C^{F}(u, v) = \frac{1}{9} \log \left(1 + \frac{(e^{9u} - 1)(e^{9v} - 1)}{e^{9} - 1} \right), (u, v) \in [0, 1]^{2}$$

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One can check it satisfies three conditions to be a copula.

Lemma: Let C' be a subcopula. Then for every $(u_1, v_1), (u_2, v_2)$ in **DomC'**,

 $|C'(u_2, v_2) - C'(u_1, v_1)| \le |u_2 - u_1| + |v_2 - v_1|.$

Hence *C* is **uniformly continuous** on its domain.

Lemma: Let C' be a subcopula. Then for any (u, v) in **Dom**C',

 $\max(u+v-1,0) \leq C'(u,v) \leq \min(u,v).$

Lemma: Let C be a copula. For any v in \mathbb{I} , the partial derivative $\partial C(u, v) / \partial u$ exist for almost all u, and for such v and u,

$$0 \leq rac{\partial}{\partial u} C(u, v) \leq 1.$$

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A distribution function is a function *F* with domain R such that
 F is nondecreasing

•
$$F(-\infty) = 0$$
 and $F(\infty) = 1$

• A joint distribution function is a function \boldsymbol{H} whose domain $\overline{\mathbb{R}}^2$ such that

• H is 2 - increasing

•
$$H(x, -\infty) = 0 = H(-\infty, y)$$
 and $H(\infty, \infty) = 1$.

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▶ The margins F and G of H are given by and .
 ▶ F(x) = H(x,∞)

• $G(y) = H(\infty, y)$

Sklar's Theorem

Let **H** be a joint distribution function with margins **F** and **G**. Then there exists a **copula C** such that for all x, y in \mathbb{R} ,

$$H(x, y) = C(F(x), G(y)).$$
(1)

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If **F** and **G** are continuous, then **C** is **unique**; otherwise, **C** is **uniquely determined** on $RanF \times RanG$.

Conversely, if C is a copula and F, G are distributions function, then the function H defined by (1) is a joint distribution function with margins F and G.



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Proof.

Lemma 1

Let H be the joint distribution function with margins F and G. Then there exists a unique subcopula C' such that

- DomC' = RanF × RanF
- For all x, y in $\overline{\mathbb{R}}, H(x, y) = C'(F(x), G(y))$.

Proof: For any $(x_1, y_1), (x_2, y_2) \in \overline{\mathbb{R}}^2$, one has

$$|H(x_2, x_2) - H(x_1, y_1)| \le |F(x_2) - F(x_1)| + |G(y_2) - G(y_1)|.$$

 $H(x_1, y_1) = H(x_2, y_2)$ when $(F(x_1), G(y_1)) = (F(x_2), G(y_2))$. Define the following mapping:

C': RanF imes RanG o [0,1]

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where C'(F(x), G(y)) := H(x, y). Hence, it is unique on $DomC' = RanF \times RanF$.

Lemma 2

Let C' be a subcopula. Then there exists a copula C such that it is coincide with C' over the DomC' and this extension is non-unique.

Proof: Because of the continuity of the subcopula C', it will be a subcopula C'' over $\overline{S_1} \times \overline{S_2}$. For any point $(a, b) \in \mathbb{I}^2$, denote by

• a_1, b_1 be the greatest element in S_1, S_2 respectively such that

$$a_1 \leq a, b_1 \leq b$$

• a_2, b_2 be the least element in S_1, S_2 respectively such that

$$a_1 \geq a, b_1 \geq b$$

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Proof.

$$egin{aligned} \lambda_1 &= egin{cases} (a-a_1)/(a_2-a_1), & ext{if } a_1 < a_2 \ 1, & ext{if } a_1 = a_2 \ \mu_1 &= egin{cases} (b-b_1)/(b_2-b_1), & ext{if } b_1 < b_2 \ 1, & ext{if } b_1 = b_2 \ \end{bmatrix} \ \mathcal{C}(a,b) &= (1-\lambda_1)(1-\mu_1)\mathcal{C}''(a_1,b_1) + (1-\lambda_1)\mu_1\mathcal{C}''(a_1,b_2) \ &+ \lambda_1(1-\mu_1)\mathcal{C}''(a_2,b_1) + \lambda_1\mu_1\mathcal{C}''(a_2,b_3) \,. \end{aligned}$$

One can check C(a, b) is grounded and 2 - increasing.

Creating a rectangle **B** in this unit square requires one more point $(c, d) \in \mathbb{I}^2$ with $c \ge a, d \ge b$.

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Let $c_1, d_1, c_2, d_2, \lambda_2, \mu_2$ in the same way with **a** and **b**.

Proof.



Figure: Scheme of points.

The **C** - volume of **B** can be decomposed into the summation of nine **C** - volumes of nine sub-rectangles, which are all non-negative.

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PART III. Copulas and Random variables.

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Lemma: Let X be continuous random variable with distribution function F, then F(X) is a standard uniform random variable, i.e., $F(X) \sim U(0, 1)$.

For any random variable X with distribution function F, the *quantile function* F^{\leftarrow} is defined by

 $F^{\leftarrow}(y) = \inf\{x \in \mathbb{R} : F(x) \ge y\}, \quad y \in [0,1],$

with the convention that $\inf \emptyset = \infty$.

Lemma: Let $X \sim U(0, 1)$ and F be any distribution function. Then $F^{\leftarrow}(X)$ is a random variable with distribution function F.

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Sklar's Theorem

Let **X** and **Y** be random variables with distribution functions **F** and **G**, respectively, and joint distribution function **H**. Then there exists a copula **C** satisfying. Then there exists a **copula C** such that for all x, y in $\overline{\mathbb{R}}$,

$$H(x, y) = C(F(x), G(y)).$$
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If **F** and **G** are continuous, then **C** is **unique**; otherwise, **C** is **uniquely determined** on $RanF \times RanG$.

• The copula in this theorem will be called the **copula of** X and Y, denoted C_{XY}

• Copula is a multivariate distribution function with standard uniform univariate margins, that is, U(0, 1) margins.

For any $(u, v) \in RanF \times RanG$, (2) can be rewritten as follows

 $H(F^{\leftarrow}(u), G^{\leftarrow}(v)) = C(u, v).$

A random vector has a continuous joint distribution function if and only if it has continuous margins.

▶ If one can estimate the copula and the margin from a given data, then the multivariate is provided.

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Example.

Let X, Y be random variables with joint distribution H

$$H(x,y) = \begin{cases} \frac{(x+1)(e^y-1)}{x+2e^y-1}, & (x,y) \in [-1,1] \times [0,\infty] \\ 1-e^{-y}, & (x,y) \in (1,\infty] \times [0,\infty] \\ 0, & \text{else where.} \end{cases}$$

with margins **F**, **G** given by

$$F(x) = \begin{cases} 0, & x < -1 \\ (x+1)/2, & x \in [-1,1], \\ 1, & x > 1 \end{cases}, G(y) = \begin{cases} 0, & y < 0 \\ 1 - e^{-y}, & y \ge 0 \end{cases}$$

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Example.

Obviously F, G are continuous functions, RanF = RanG = I and the quantile functions of F, G are

$$F^{\leftarrow}(x) = 2x - 1,$$
 $G^{\leftarrow}(y) = -\ln(1 - \nu).$

Because RanF = RanG = I, one can find the copula

$$C(u,v)=\frac{uv}{u+v-uv}.$$

Hence,

$$H(x,y) = \frac{F(x)G(y)}{F(x) + G(y) - F(x)G(y)}$$

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Independent Copula

Theorem

Let **X** and **Y** be continuous random variable. Then **X** and **Y** are independent if and only if $C_{XY} = \prod$



Figure: Graph of product copula and its contour graph.

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Invariance principle.

Theorem.

Let **X** and **Y** are continuous random variables with distribution functions **F** and **G** respectively and a copula C_{XY} . If α, β are strictly increasing on $RanX \times RanY$, respectively, then

 $C_{XY} = C_{\alpha(X)\beta(Y)}.$

 \blacktriangleright If α is strictly increasing and β is strictly decreasing, then

$$C_{\alpha(X),\beta(Y)}(u,v) = u - C_{XY}(u,1-v)$$

 \blacktriangleright If α is strictly decreasing and β is strictly increasing, then

$$C_{\alpha(X),\beta(Y)}(u,v)=v-C_{XY}(1-u,v)$$

• If α and β are both strictly decreasing, then

$$C_{\alpha(X),\beta(Y)}(u,v) = u + v - 1 + C_{XY}(1-u,1-v)$$

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Theorem.

If **X** and **Y** are two random variables with **F**, **G** are their distribution function respectively. Then for any x, y in $\overline{\mathbb{R}}$, and any u, v in [0, 1], one has

 $\max(u+v-1,0) \leq C(u,v) \leq \min(u,v)$

 $\max(F(x) + G(y) - 1, 0) \leq H(x, y) \leq \min(F(x), G(y))$

► Call $W(u, v) = \max(u + v - 1, 0)$, $M(u, v) = \min(u, v)$. Then W(u, v) and M(u, v) are copulas

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The Fréchet-Hoeffding Bounds



Figure: The Fréchet-Hoeffding Bounds Lower bound and Upper bound.

Image: A matrix

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Example

Considering the Frank Copula

$$C^{F}(u,v) = rac{1}{9} \log \left(1 + rac{(e^{9u} - 1)(e^{9v} - 1)}{e^{9} - 1}
ight), (u,v) \in [0,1]^{2}$$



Figure: Frank copula and its density.

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Example



Figure: n = 1000 independent observations $(U_1, U_2) \sim C^F$.

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Natural Question: What happens when the equality occurs?

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A subset **S** of \mathbb{R}^2 is nondecreasing if for any (x, y), (u, v) in **S**, x < u implies $y \le v$.



Figure: An example of nondecreasing set.

A subset **S** of \mathbb{R}^2 is nonincreasing if for any (x, y), (u, v) in **S**, x < u implies $y \ge v$.

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▶ The **support** of a copula *C* is the complement of the union of all open subset of \mathbb{I}^2 with *C* - measure zero.

Theorem.

Let **X** and **Y** be random variables with joint distribution function **H**. Then **H** is equal to its **Fréchet-Hoeffding upper bound** if and only if the support of **H** is a **nondecreasing subset** of \mathbb{R}^2 .

Analogously, H is equal to its **Fréchet-Hoeffding lower bound** if and only if the support of H is a **nonincreasing subset** of \mathbb{R}^2 .

Example: The Fréchet-Hoeffding upper bound is a copula and its support set is the main diagonal.

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