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![](_page_0_Picture_1.jpeg)

![](_page_0_Picture_2.jpeg)

Change-point methods for anomaly detection in fibrous media

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### Fibre materials Random sequential adsorption (RSA)

![](_page_1_Figure_3.jpeg)

Visualization of layered RSA image with  $200 \times 200 \times 300$  voxels (left) and homogeneous RSA image with  $500 \times 500 \times 500$  voxels (right), radius 4

## **Problem setting**

- A fibre  $\gamma$  is a simple curve  $\{\gamma(t), t \in [0, 1]\}$  in  $\mathbb{R}^3$  of finite length.

- The collection of fibres forms a fibre system  $\phi$ .
- The length measure  $\phi(B) = \sum_{\gamma \in \phi} h(\gamma \cap B)$ , where *h* is the length of fibre in window  $B \subset \mathbb{R}^3$ .

- A fibre process  $\Phi$  is a random element with values in the set  $\mathbb{D}$  of all fibre systems  $\phi$  with  $\sigma$ -algebra D generated by sets of the form  $\{\phi \in \mathbb{D} : \phi(B) < x\}$ .

## Classification

- Let w(x) be some characteristic of a fibre at point x: fibre local direction, curvature, etc.

- A weighted random measure

 $\Psi(B \times L) = \int_B \mathbb{1}\{w(x) \in L\}\Phi(dx).$ 

- If the fibre process  $\Phi$  is stationary, then

 $\mathbf{E}\Psi(\mathbf{B} \times \mathbf{L}) = \lambda |\mathbf{B}| f(\mathbf{L})$ , where  $\lambda$  is called the intensity of  $\Psi$ ,

- A probability measure f on  $\mathbb{S}^2$  is called the directional distribution of fibres.

## Testing

 $H_0: \Phi$  is stationary with intensity  $\lambda$  and directional distribution f vs.

 $H_1$  : There exists a compact set  $A \subset W$  with |A| > 0 and  $|W \setminus A| > 0$  such that

$$\frac{1}{\lambda|A|}\mathsf{E}\int_{A}\mathbb{1}\{w(x)\in\cdot\}\Phi(dx)\neq\frac{1}{\lambda|W\setminus A|}\mathsf{E}\int_{W\setminus A}\mathbb{1}\{w(x)\in\cdot\}\Phi(dx).$$

If  $H_1$  holds true, the region A is called an anomaly region.

#### Data

- The dilated fibre system  $\Phi \oplus B_r \cap W$  in window  $W \subset \mathbb{R}^3$  is observed as a 3D greyscale image.

- Reconstruction of  $\mu$ CT image by MAVI: Modular Algorithms for volume Images, provided by Fraunhofer ITWM.

![](_page_5_Picture_4.jpeg)

![](_page_5_Picture_5.jpeg)

#### Local directions and clustering criteria

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- A separation of fibres (and estimation of their directions) requires large computational resources for 3D images.

- An estimation of local direction is much faster but produces dependent sample.

- In each  $W_l$ , the "average local direction" is computed using SubfieldFibreDirection in MAVI.

- We group  $W_l$  in classification windows  $W_l$ .

- For each  $W_l$  we assign a classification attribute: entropy, average direction.

## **Entropy estimation**

- Entropy of an absolutely continuous  $\mathbb{S}^2$ -valued random variable with density *f* is  $E_X = -\int_{\mathbb{S}^2} \log(f(x))f(x)\sigma(dx)$ , where  $\sigma$  is the spherical surface measure on  $\mathbb{S}^2$ .

- Plug-in estimators required large samples. In simulations for uniform distribution on a sphere  $N > 50^3$ .
- Nearest neighbour estimator (KL-estimator, 1987)

$$\hat{E} = \frac{d}{N} \sum_{i=1}^{N} \log \rho_i + \log(c(N-1)) + \gamma.$$

- We propose modification of the nearest neighbour estimator:

$$\hat{E}_M = \frac{d}{N_M} \sum_{i=1}^N \mathbb{1}\{\rho_i > \rho_0\} \log \rho_i + \log(c(N_M - 1)) + \gamma,$$

where  $N_M = \sum_{i=1}^N \mathbb{1}\{\rho_i > \rho_0\}, \rho_0$  is a penalty value.

## Entropy estimation: homogeneous RSA (random sequential adsorption) image

![](_page_8_Picture_2.jpeg)

![](_page_8_Figure_3.jpeg)

Histogram of frequencies of the local entropy

 $2000\times2000\times2100$  voxels

![](_page_9_Picture_0.jpeg)

#### Entropy estimation: layered RSA image

![](_page_9_Picture_3.jpeg)

 $2000\times2000\times2100$  voxels

![](_page_9_Figure_5.jpeg)

# Histogram of frequencies of the local entropy

![](_page_9_Picture_7.jpeg)

## Random fields with inhomogeneities in mean

- Let be  $\{\xi_k, k \in \mathbb{Z}^3\}$  an integrable, centered, stationary, real-valued random field.

- { $\xi_k, k \in \mathbb{Z}^3$ } is *m*-dependent, and there exist  $H, \sigma > 0$  such that  $\mathbf{E}|\xi_k|^p \leq \frac{p!}{2}H^{p-2}\sigma^2, p = 2, 3, \dots$ 

- Let  $\Theta$  be a finite parametric space. For every  $\theta \in \Theta$  we define subspace of anomalies  $I_{\theta} \subset \mathbb{Z}^3$ .

- For some  $\theta_0\in\Theta$  we observe

$$\mathbf{s}_{\mathbf{k}} = \xi_{\mathbf{k}} + \mu + h\mathbb{1}\{\mathbf{k} \in \mathbf{I}_{\theta_0}\}, \mathbf{k} \in \mathbf{W}.$$

- Let  $\Theta_0$  correspond to the significant anomalies, i.e, for  $\gamma_0, \gamma_1 \in (0, 1)$  we let  $\Theta_0 = \{\theta \in \Theta : \gamma_0 | W | \le |I_{\theta}| \le \gamma_1 | W |\}$ . - Then  $\Theta_1 = \Theta \setminus \Theta_0$  corresponds to the extremely small or large anomalies, i.e,

 $\Theta_1 = \{ \theta \in \Theta : |I_{\theta}| < \gamma_0 |W|, \text{ or } |I_{\theta}| > \gamma_1 |W| \}.$ 

#### Testing the change of expectation

- The change-point hypotheses for the random field  $\{s_k, k \in W\}$  with respect to its expectation

$$H_0$$
: **E** $s_k = \mu$  for every  $k \in W$  (i.e.  $h = 0$ ) vs.

$$H_1$$
: There exists  $\theta_0 \in \Theta_0$  such that

$$\mathbf{E}\mathbf{s}_{k} = \mu + h, k \in I_{\theta_{0}}, \ h \neq 0, \text{ and } \mathbf{E}\mathbf{s}_{k} = \mu, k \in I_{\theta_{0}^{c}}.$$

- Analogue of CUSUM statistics

$$\begin{split} Z(\theta) &= \frac{1}{|I_{\theta}|} \sum_{k \in I_{\theta}} s_k - \frac{1}{|I_{\theta}^c|} \sum_{k \in I_{\theta}^c} s_k \\ &= \frac{1}{|I_{\theta}|} \sum_{k \in I_{\theta}} \xi_k - \frac{1}{|I_{\theta}^c|} \sum_{k \in I_{\theta}^c} \xi_k + h\left(\frac{|I_{\theta} \cap I_{\theta_0}|}{|I_{\theta}|} - \frac{|I_{\theta}^c \cap I_{\theta_0}|}{|I_{\theta}^c|}\right). \end{split}$$

- Test statistics:  $T_W = \max_{\theta \in \Theta_0} |Z(\theta)|$ .

#### **Test statistics**

- Critical values  $y_{\alpha}$  via the probability of the 1st-type error:

$$\mathbf{P}_{H_0}(T_W \ge y_\alpha) = \mathbf{P}\left(\max_{\theta \in \Theta_0} \left| \frac{1}{|I_\theta|} \sum_{k \in I_\theta} \xi_k - \frac{1}{|I_\theta^c|} \sum_{k \in I_\theta^c} \xi_k \right| \ge y_\alpha\right) \le \alpha.$$

- Tail probabilities

$$\begin{aligned} \mathbf{P}_{H_0}(T_W \ge \mathbf{y}) &\leq \sum_{\theta \in \Theta_0: |I_{\theta}^c| \le \frac{\sigma^2 |W|}{\mathbf{y} H}} 2 \exp\left(-\frac{\mathbf{y}^2}{4m^3 \sigma^2} \frac{|I_{\theta}^c| |I_{\theta}|}{|W|}\right) \\ &+ \sum_{\theta \in \Theta_0: |I_{\theta}^c| > \frac{\sigma^2 |W|}{\mathbf{y} H}} 2 \exp\left(-\frac{\mathbf{y}}{2Hm^3} |I_{\theta}| + \frac{\sigma^2 |W|}{4H^2m^3 |I_{\theta}^c|} |I_{\theta}|\right). \end{aligned}$$

## **Tail probabilities**

- If  $\xi_k$ 's are Gaussian, then  $H = \sigma$ . If  $|\xi_k| \leq M$  then H = M.
- Particularly, if  $y < \frac{\sigma^2}{H(1-\gamma_0)}$  then

$$\mathbf{P}_{\mathcal{H}_0}(\mathcal{T}_W \geq \mathbf{y}) \leq 2 \left|\Theta_0\right| \exp\left(-\frac{\mathbf{y}^2}{4m^3\sigma^2} |W|\gamma_0(1-\gamma_0)\right),$$

and if  $y > \frac{\sigma^2}{H(1-\gamma_1)}$  then

$$\mathbf{P}_{H_0}(T_W \ge y) \le 2 |\Theta_0| \exp\left(-\frac{y}{4Hm^3}\gamma_0|W|\right).$$

### Simulated data

## Homogeneous RSA data:

Attr.	$ \Theta_0 $	Var.	Test stat.	<i>p</i> -value	<sup>14</sup>
ĩ	39395	0.04360	0.0344	1.00	1.0
ỹ	39395	0.03743	0.0130	1.00	0.8
ĩ	39395	0.03749	0.0146	1.00	0.4
Ĩ	16536	0.08984	0.0942	1.00	0.0 -1.5 -1.0 -0.5 0.0 0.5

#### Layered RSA data:

Attr.	$ \Theta_0 $	Var.	Test stat.	p-value 🔐	
ĩ	39395	0.10592	0.44036	$4.6  imes 10^{-30}$	
ỹ	39395	0.10948	0.43163	$2.8 imes10^{-23}$ $_{\circ4}$	
ĩ	39395	0.06151	0.18764	0.3012	
Ĩ	16536	0.3583	1.07030	0.00	5 -1.0 -0.5 0.0 0.5 1.0

Classification by spatial SAEM algorithm: layered RSA image.

![](_page_15_Figure_2.jpeg)

 $2000\times2000\times2100$  voxels

Spatial SAEM classification by entropy and mean local directions

# Real glass fibre reinforced polymer

- The images are provided by the Institute for Composite Materials (IVW) in Kaiserslautern: 970 × 1469 × 1217 voxels, the estimated radius of 3 voxels. We obtain 64 × 97 × 80 small windows  $\widetilde{W}_i$  with 15 × 15 × 15 voxels.

- Change point testing:

Attr.	Н	$\sigma^2$	т	$ \Theta_0 $	Var.	Test stat.	<i>p</i> -value
ĩ	0.5	0.2	7	33004	0.04589	0.15995	1.00
ŷ	0.5	0.2	7	33004	0.06795	0.44733	$2.1  imes 10^{-10}$
ĩ	0.5	0.2	7	33004	0.07982	0.43383	$1.3 imes10^{-6}$
Е	0.7071	0.5	1	12366	0.30126	0.46811	$3.96 imes10^{-8}$

#### Classification by spatial SAEM algorithm: real data.

![](_page_17_Figure_2.jpeg)

![](_page_17_Figure_3.jpeg)

 $970\times1469\times1217$  voxels

Spatial SAEM classification by entropy and mean local directions

![](_page_18_Picture_0.jpeg)

Preprint: Detecting anomalies in fibre systems using 3-dimensional image data. arXiv:1907.06988.

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#### Thank you for your attention!