



Probabilistic prediction of solar power supply to distribution networks, using global radiation forecasts

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Outline

1 General context

- Risk in feed-in of solar power
- Visualization of data
- Modeling idea

2 Bivariate copulas

- Archimedean copulas
- Fitting process
- Results

3 Vine copulas

- D-vine copulas
- Fitting process
- Results

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Risk in feed-in of solar power

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- Increase in solar plants → voltage violations and overloading problems

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Timeframe: May, June and July of the years 2015-2017 (11-12 UTC)

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Goal

Predict the risk of solar power supply exceeding critical thresholds

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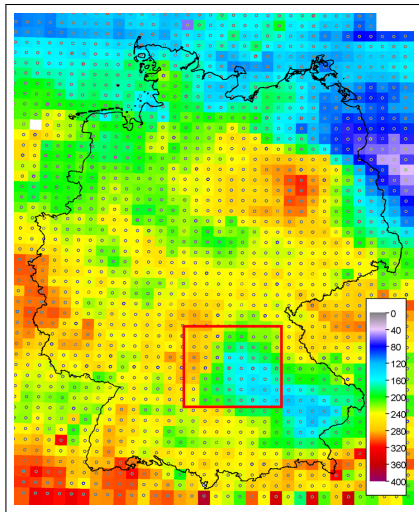
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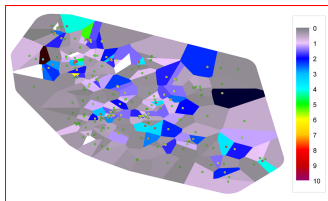
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Visualization of data

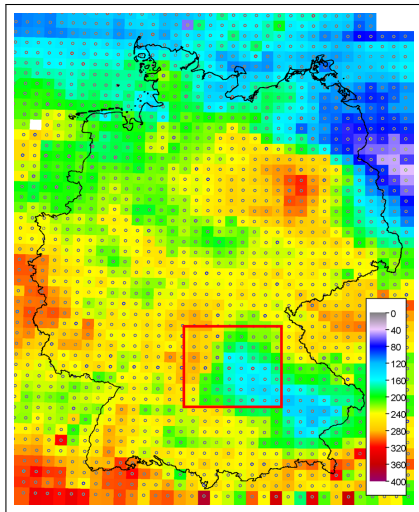


Global radiation forecast (in J/cm^2)
for July 07, 2017 11-12 UTC

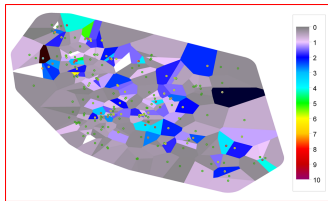


Measured solar power supply (in MW)
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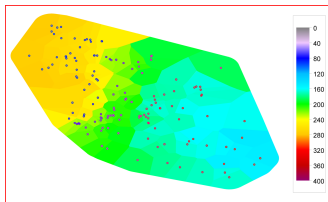
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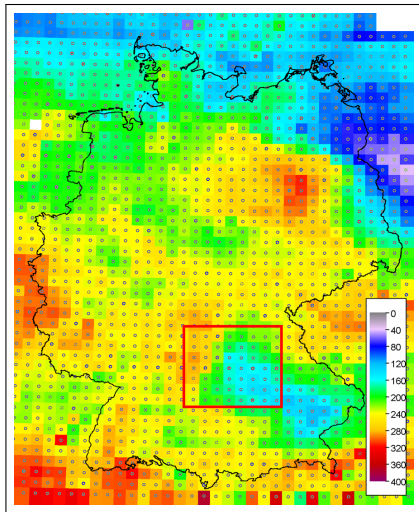


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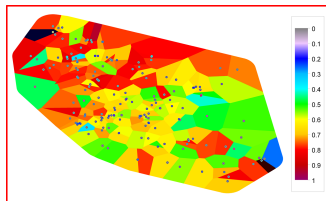


Interpolated global radiation forecast
for July 07, 2017 11-12 UTC

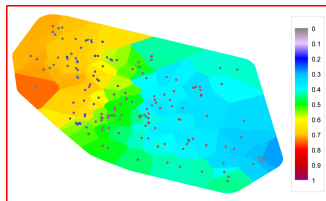
Visualization of data



Global radiation forecast (in J/cm^2)
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Normalized solar power supply
for July 07, 2017 11-12 UTC



Normalized global radiation forecast
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Copula models

Random variables

R : (Normalized) global radiation forecast

S : (Normalized) solar power supply

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Goals

For a predefined threshold v and feed-in points p_1, \dots, p_n compute the conditional probabilities

- $P(S_1 \geq v \mid R_1 = r(p_1, t))$
- $P(S_1 + \dots + S_n \geq v \mid R_1 = r(p_1, t), \dots, R_n = r(p_n, t))$

given global radiation forecasts $r(p_1, t), \dots, r(p_n, t)$ and forecast time t

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Modeling approach

- Fit univariate marginal distributions
- Fit **bivariate** and **multivariate** distributions using **bivariate copulas** and **D-vine copulas**

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Copula theory

Bivariate copulas

A **bivariate copula** is the joint distribution function $C : [0, 1] \times [0, 1] \rightarrow [0, 1]$ of a 2-dimensional random vector (U, V) with components U and V **uniformly distributed** on $[0, 1]$

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Theorem of Sklar

Let (R, S) be a 2-dimensional random vector with joint distribution function $F_{(R,S)} : \mathbb{R}^2 \rightarrow [0, 1]$ and marginal distribution functions F_R and F_S . Then, a **bivariate copula** function $C : [0, 1] \times [0, 1] \rightarrow [0, 1]$ **exists** such that

$$F_{(R,S)}(r, s) = C(F_R(r), F_S(s)) \quad \text{for all } r, s \in \mathbb{R}$$

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Differential form of Sklar's theorem

For the **density functions** $f_{(R,S)}$, f_R , f_S and c it holds that

$$f_{(R,S)}(r, s) = f_R(r) \cdot f_S(s) \cdot c(F_R(r), F_S(s)) \quad \text{for all } r, s \in \mathbb{R}$$

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Definition of Archimedean copulas

Archimedean generator

A function $g : [0, 1] \rightarrow [0, \infty]$ is called **Archimedean generator** if g is continuous, strictly decreasing and solves $g(1) = 0$.

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Pseudo-inverse

The **pseudo-inverse** $g^{[-1]}$ of an Archimedean generator g is an extension of the inverse function $g^{(-1)}$ defined as

$$g^{[-1]}(t) = \begin{cases} g^{(-1)}(t), & \text{if } 0 \leq t \leq g(0) \\ 0, & \text{if } g(0) \leq t \leq \infty. \end{cases}$$

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Archimedean copula

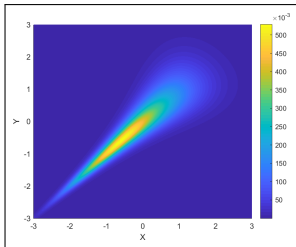
The **Archimedean copula** generated by g is given by

$$C(u, v) = g^{[-1]}(g(u) + g(v)) \quad u, v \in [0, 1].$$

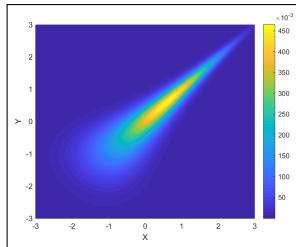
Examples of Archimedean copulas

copula family	Archimedean generator	parameter
Joe	$g(t) = -\log(1 - (1 - t)^\theta)$	$\theta \in [1, \infty]$
Frank	$g(t) = (-\log(\frac{\exp(-\theta t) - 1}{\exp(-\theta) - 1}))^\theta$	$\theta \in \mathbb{R} \setminus \{0\}$
Clayton	$g(t) = \frac{1}{\theta}(t^{-\theta} - 1)$	$\theta \in [-1, \infty) \setminus \{0\}$
Gumbel	$g(t) = (-\log(t))^\theta$	$\theta \in [1, \infty)$

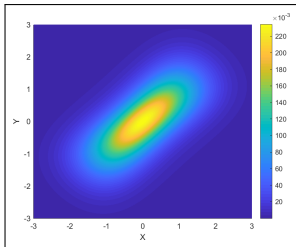
Visualization of the copula types



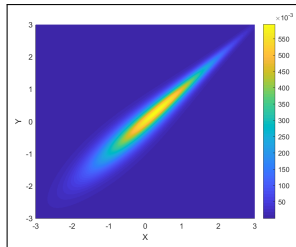
Clayton with parameter $\theta = 5$



Joe with parameter $\theta = 5$



Frank with parameter $\theta = 5$



Gumbel with parameter $\theta = 5$

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Fitting of univariate marginal distributions

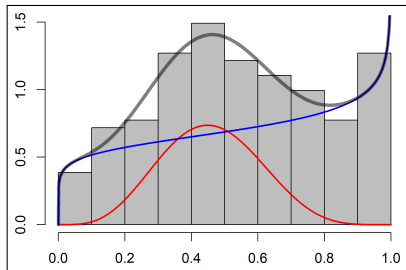
- Fit **mixed beta densities** $f(x) = qf_1(x) + (1 - q)f_2(x)$ with mixture parameter $q \in [0, 1]$ and beta densities $f_i : (0, 1) \rightarrow [0, \infty)$ with $f_i(x) = \frac{\Gamma(a_i+b_i)}{\Gamma(a_i)\Gamma(b_i)} x^{a_i-1}(1-x)^{b_i-1}$ and two parameters $a_i, b_i > 0$

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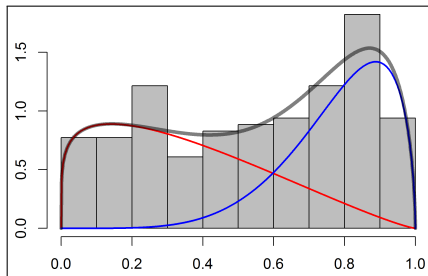
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- Apply **EM algorithm** to estimate the parameters of the mixed beta densities

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Global radiation forecast data



Solar power supply data

Fitting bivariate copulas

- Estimate the copula parameter θ for each copula type by **maximizing the likelihood function** with given F_R and F_S

Fitting bivariate copulas

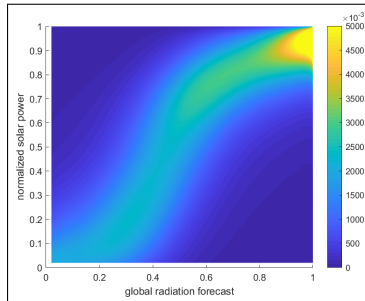
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Fitted bivariate joint density

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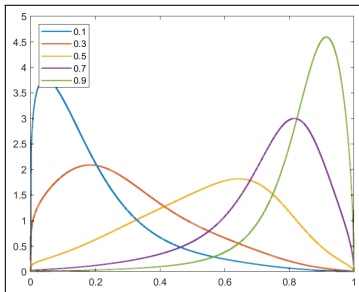
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Computation of conditional probabilities

Computation of conditional probabilities

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$$f_S(s | R = r) = \frac{f_{(R,S)}(r,s)}{f_R(r)} = f_S(s) \cdot c(F_R(r), F_S(s))$$



Conditional densities given the global radiation forecasts r for $r = 0.1, 0.3, 0.5, 0.7$ and 0.9

Computation of conditional probabilities

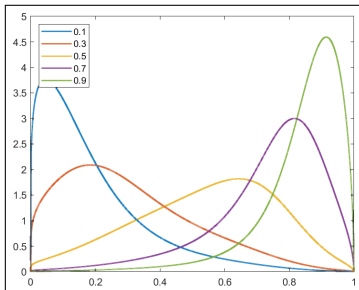
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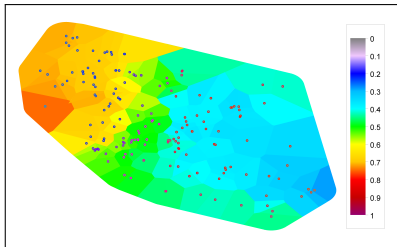
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$$P(v, r) = P(S \geq v | R = r) = \int_v^1 f_S(s | R = r) ds$$

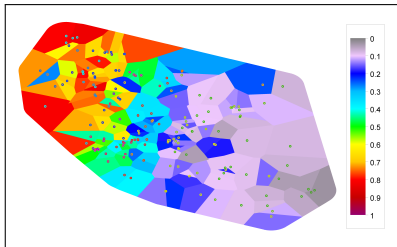


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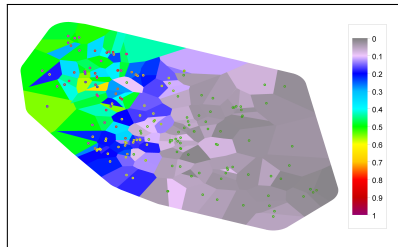
Conditional probabilities for a critical event



Normalized global radiation forecast

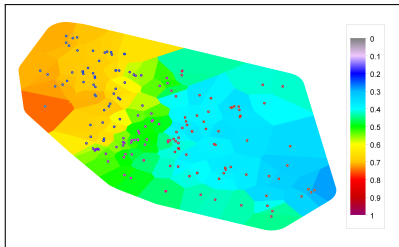


Conditional probabilities for threshold $\nu = 0.7$

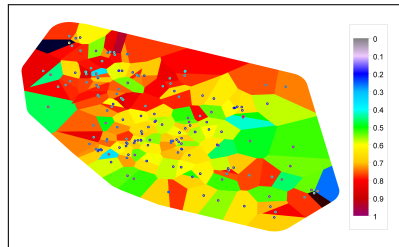


Conditional probabilities for threshold $\nu = 0.8$

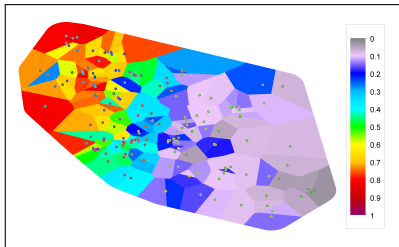
Conditional probabilities for a critical event



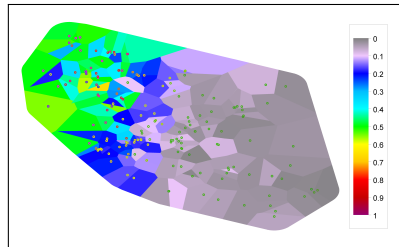
Normalized global radiation forecast



Normalized solar power supply



Conditional probabilities for threshold $\nu = 0.7$



Conditional probabilities for threshold $\nu = 0.8$

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Vine copula

A vine copula is obtained by **decomposing a multivariate density** into conditional densities and applying **Sklar's theorem sequentially** to each conditional density

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Vine copulas as families of trees

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- each **edge** is a conditional bivariate copula

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Vine copulas as families of trees

We can interpret a vine copula as a **family of trees**, where

- each **edge** is a conditional bivariate copula
- each **node** is a conditional cumulative distribution function

R-vine

A **regular vine** (short: R-vine) V on n elements is a family of trees

$\{\mathcal{T}_1, \dots, \mathcal{T}_{n-1}\}$ with edges $E(V) = E_1 \cup \dots \cup E_{n-1}$, such that

1. $\mathcal{T}_1 = (N_1, E_1)$ is a connected tree with nodes $N_1 = \{1, \dots, n\}$ and edges E_1
2. \mathcal{T}_k is a tree with nodes $N_k = E_{k-1}$ for all $k \in \{2, \dots, n-1\}$
3. $\#(e_1 \Delta e_2) = 2$ for all $\{e_1, e_2\} \in E_k$ with $k \in \{2, \dots, n-1\}$

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R-vine decomposition

The **decomposition** of a n -dim. density $f_{1, \dots, n}$ corresponding to an R-vine V with edges $E(V)$ is given by

$$f_{1, \dots, n} = \prod_{e \in E(V)} c_{t_1, t_2 | S(e)}(F_{t_1 | S(e)}, F_{t_2 | S(e)}) \cdot \prod_{j=1}^n f_j,$$

where $S(e)$ is the so-called **conditioning set**, $T(e) = \{t_1, t_2\}$ is the **conditioned set** of the edge e and f_j are the one-dim. marginal densities

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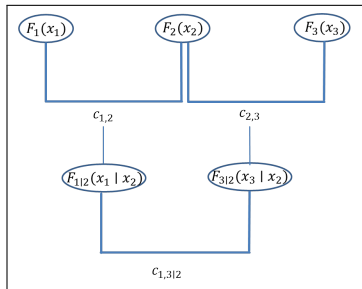
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D-vine copulas

D-vines

- D-vines are a special **type** of R-vine

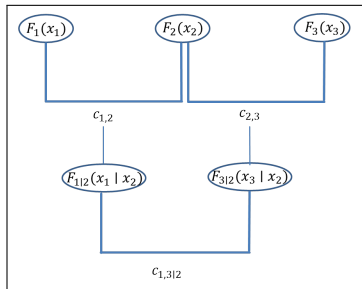


D-vine structure for 3 random variables

D-vine copulas

D-vines

- D-vines are a special **type** of R-vine
- Each node is connected to **not more than 2 edges**

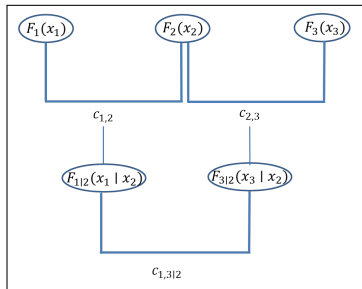


D-vine structure for 3 random variables

D-vine copulas

D-vines

- D-vines are a special **type** of R-vine
- Each node is connected to **not more than 2 edges**
- For each edge in $E_k = \{e_1, \dots, e_{n-k}\}$ the **conditioned set** is $T(e_i) = \{i, i+k\}$ and the **conditioning set** is $S(e_i) = \{i+1, \dots, i+k-1\}$



D-vine structure for 3 random variables

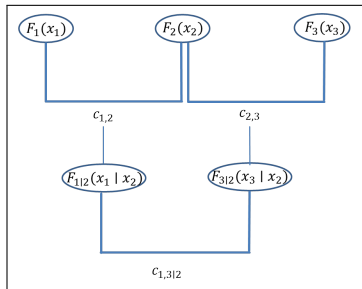
D-vine copulas

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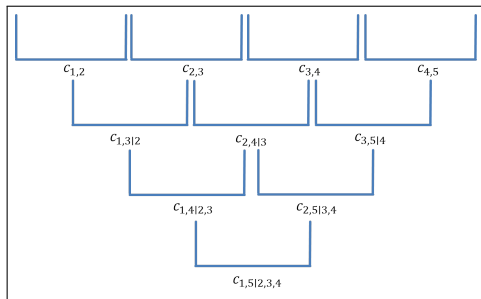
For 3-dim. densities a D-vine corresponds to following **decomposition**:

$$\begin{aligned}
 f_{1,2,3}(x_1, x_2, x_3) &= f_{3|1,2}(x_3 | x_1, x_2) f_{2|1}(x_2 | x_1) f_1(x_1) \\
 &= c_{1,3|2}(F_{1|2}(x_1 | x_2), F_{3|2}(x_3 | x_2)) f_{3|2}(x_3 | x_2) f_{2|1}(x_2 | x_1) f_1(x_1) \\
 &= c_{1,3|2}(F_{1|2}(x_1 | x_2), F_{3|2}(x_3 | x_2)) c_{2,3}(F_2(x_2), F_3(x_3)) f_3(x_3) \\
 &\quad c_{1,2}(F_1(x_1), F_2(x_2)) f_2(x_2) f_1(x_1)
 \end{aligned}$$



D-vine structure for 3 random variables

D-vine copulas



D-vine structure for 5 random variables

The decomposition of an n -dim. density corresponding to D-vines is:

$$f_{1,\dots,n}(x_1, \dots, x_n) = \prod_{k=1}^{n-1} \prod_{i=1}^{n-k} c_{i,i+k|i+1,\dots,i+k-1}(F_{i|i+1,\dots,i+k-1}(x_i | x_{i+1}, \dots, x_{i+k-1})),$$

$$F_{i+k|i+1,\dots,i+k-1}(x_{i+k} | x_{i+1}, \dots, x_{i+k-1})) \cdot \prod_{j=1}^n f_j(x_j)$$

Outline

- 1 General context
 - Risk in feed-in of solar power
 - Visualization of data
 - Modeling idea

- 2 Bivariate copulas
 - Archimedean copulas
 - Fitting process
 - Results

- 3 Vine copulas
 - D-vine copulas
 - **Fitting process**
 - Results

Fitting of vine copulas

Sequential estimation

The following steps are applied, starting with the first row:

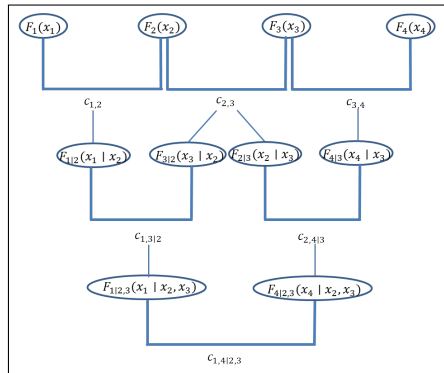
Step 0: Fit the marginal cdfs

Step 1: Transform the data based on the computed cdfs

Step 2: Fit bivariate copulas to the transformed data

Step 3: Compute conditional cdfs using the bivariate copulas

Step 4: Repeat Step 1-3 till the end



Fitting a D-vine with 4 random variables

Fitting of vine copulas

Sequential estimation

The following steps are applied, starting with the first row:

Step 0: Fit the marginal cdfs

Step 1: Transform the data based on the computed cdfs

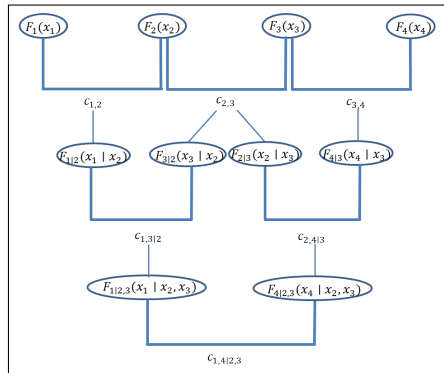
Step 2: Fit bivariate copulas to the transformed data

Step 3: Compute conditional cdfs using the bivariate copulas

Step 4: Repeat Step 1-3 till the end

Fitting of bivariate copulas

Apply **ML estimation** to fit one-parametric Archimedean copulas



Fitting a D-vine with 4 random variables

Outline

- 1 General context
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Compute conditional probabilities

Goal

For a predefined threshold v and feed-in points p_1, \dots, p_n compute the **conditional probabilities** $P(S_1 + \dots + S_n \geq v \mid R_1 = r(p_1, t), \dots, R_n = r(p_n, t))$ given global radiation forecasts $r(p_1, t), \dots, r(p_n, t)$ and forecast time t

Application of D-vine copulas

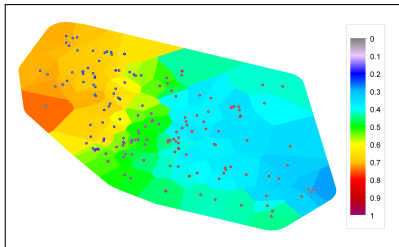
Fit an $n + 1$ -dim. **D-vine copula** to the random vector $(R_1, \dots, R_n, S_1 + \dots + S_n)$

Computation of conditional probabilities

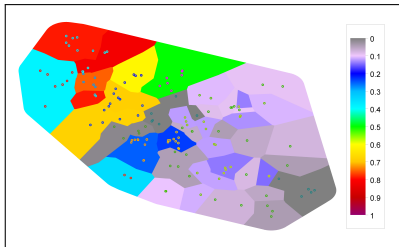
Based on the **fitted D-vine copula** we compute

$$P(S_1 + \dots + S_n \geq v \mid R_1 = r(p_1, t), \dots, R_n = r(p_n, t)) = \int_v^1 c_{1,n+1|2,\dots,n}(F_{1|2,\dots,n}(r(p_1, t) \mid r(p_2, t), \dots, r(p_n, t)), F_{n+1|2,\dots,n}(s \mid r(p_1, t), \dots, r(p_{n-1}, t))) ds$$

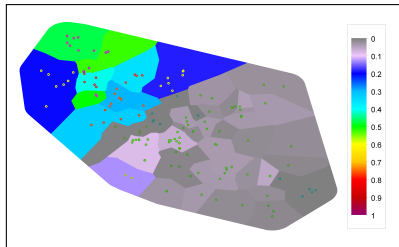
Conditional probabilities calculated by multivariate D-vines



Normalized global radiation forecast

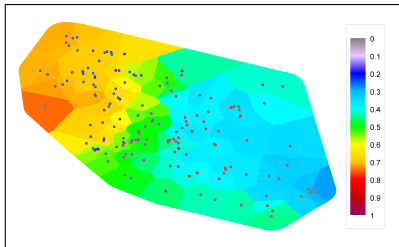


Conditional probabilities for threshold $\nu = 0.7$

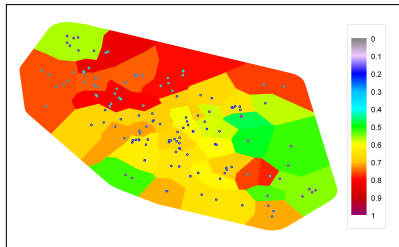


Conditional probabilities for threshold $\nu = 0.8$

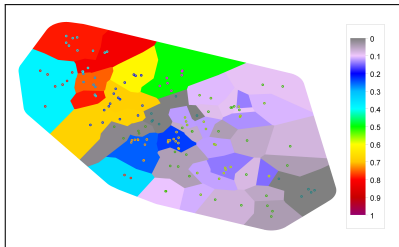
Conditional probabilities calculated by multivariate D-vines



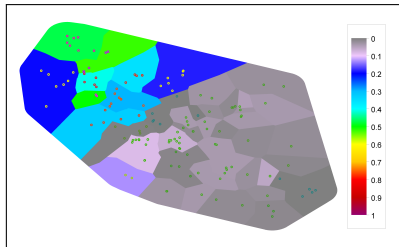
Normalized global radiation forecast



Normalized aggregated solar power supply








Conditional probabilities for threshold $v = 0.7$



Conditional probabilities for threshold $v = 0.8$

Literature

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-  Dempster, A. P., Laird, N. M., and Rubin, D. B., 1977. Maximum likelihood from incomplete data via the EM algorithm. *Journal of the Royal Statistical Society: Series B (Methodological)*, 39(1), 1-22.
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-  Joe, H., 2014. *Dependence Modeling with Copulas*. Chapman and Hall/CRC.
-  von Loeper, F., Schaumann, P., de Langlard, M., Hess, R., Bäsmann, R. and Schmidt, V., 2019. Probabilistic prediction of solar power supply to distribution networks, using forecasts of global radiation. Preprint (submitted)