



## Long range dependence for heavy tailed random functions

Joint work with R. Kulik (U Ottawa),  
V. Makogin, M. Oesting (U Siegen), A.  
Rapp

## Overview

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## Introduction: Random functions with long memory

Random function = Set of random variables indexed by  $t \in T$ .

Let  $X = \{X_t, t \in T\}$  be a wide sense stationary random function defined on an abstract probability space  $(\Omega, \mathcal{F}, P)$ , e.g.,  $T \subseteq \mathbb{R}^d$ ,  $d \geq 1$ . The property of **long range dependence (LRD)** can be defined as

$$\int_T |C(t)| dt = +\infty$$

where  $C(t) = \text{cov}(X_0, X_t)$ ,  $t \in T$  (McLeod, Hipel (1978); Parzen (1981)). Sometimes one requires that  $C \in RV(-a)$ , i.e.,  $\exists a \in (0, d)$  such that

$$C(t) = \frac{L(t)}{|t|^a}, \quad |t| \rightarrow +\infty,$$

where  $L(\cdot)$  is a slowly varying function.

## Various approaches to define LRD

- ▶ Unbounded **spectral density** at zero.
- ▶ Growth order of **sums' variance** going to infinity.
- ▶ **Phase transition** in certain parameters of the function (stability index, Hurst index, heaviness of the tails, etc.) regarding the different **limiting behaviour of some statistics** such as
  - ▶ **Partial sums**
  - ▶ **Partial maxima.**

These approaches are not equivalent, often **statistically not tractable** and **tailored for a particular class** of random functions (e.g., time series, square integrable, stable, etc.)

## Various approaches to define LRD

LRD for heavy tailed random functions:

- ▶ Phase transitions in the limiting behaviour of **partial sums and maxima** of inf. divisible random processes and their ergodic properties (Samorodnitsky 2004, Samorodnitsky & Roy 2008, Roy 2010).
- ▶  **$\alpha$ -spectral covariance approach** for linear random fields with innovations lying in the domain of attraction of  $\alpha$ -stable law (Paulauskas (2016), Damarackas, Paulauskas (2017))

## LRD: Infinite variance case

For a stationary random function  $X$  with  $E X_t^2 = +\infty$  introduce

$$\text{cov}_X(t, u, v) = \text{cov}(\mathbb{1}(X_0 > u), \mathbb{1}(X_t > v)), \quad t \in T, u, v \in \mathbb{R}.$$

It is always defined as the indicators involved are bounded functions.

A random function  $X$  is called **SRD** (**LRD**, resp.) if

$$\sigma_{\mu, X}^2 = \int_T \int_{\mathbb{R}^2} |\text{cov}_X(t, u, v)| \mu(du) \mu(dv) dt < +\infty \quad (= +\infty)$$

for **all** finite measures (for **a** finite measure, resp.)  $\mu$  on  $\mathbb{R}$ . For discrete parameter random fields (say, if  $T \subseteq \mathbb{Z}^d$ ), the  $\int_T dt$  in the above line should be replaced by a  $\sum_{t \in T: t \neq 0}$ .

## Motivation

Assume that  $X$  is wide sense stationary with covariance function  $C(t) = \text{cov}(X_0, X_t)$ ,  $t \in T$ , and moreover,

$$\text{cov}_X(t, u, v) \geq 0 \text{ or } \leq 0 \text{ for all } t \in T, u, v \in \mathbb{R}.$$

Examples of  $X$  with this property are all **PA** or **NA**- random functions. **W. Hoeffding (1940)** proved that

$$C(t) = \int_{\mathbb{R}^2} \text{cov}_X(t, u, v) du dv. \quad (1)$$

Then,  $X$  is long range dependent if

$$\int_T |C(t)| dt = \int_T \int_{\mathbb{R}^2} |\text{cov}_X(t, u, v)| du dv dt = +\infty.$$

## Motivation: memory and excursions

Level (excursion) sets and their volumes:

Let  $a_n(u) = \nu_d(A_u(X, W_n))$  be the volume of the excursion set

$$A_u(X, W_n) = \{t \in T \cap W_n : X_t > u\}$$

of a random field  $X$  at level  $u$  in an observation window  $W_n = n \cdot W$  where  $W \subset \mathbb{R}^d$  is a convex body.



## Motivation: excursions and SRD

Multivariate CLT for level sets' volumes (Bulinski, S., Timmermann, Karcher, 2012):

For a stationary centered weakly dependent random function  $X$  satisfying some additional conditions (square integrable,  $\alpha$ - or max-stable, inf. divisible) we have for any levels  $u, v \in \mathbb{R}$  that

$$\frac{(a_n(u), a_n(v))^{\top} - (\mathbb{P}(X_0 \geq u), \mathbb{P}(X_0 \geq v))^{\top} \cdot \nu_d(W_n)}{\sqrt{\nu_d(W_n)}} \xrightarrow{d} \mathcal{N}(\mathbf{0}, \Sigma)$$

as  $n \rightarrow \infty$ . Here  $\Sigma = (\sigma_{ij})_{i,j=1}^2$  with  $\sigma_{12} = \int_{\mathbb{R}^d} \text{cov}_X(t, u, v) dt$ .

So,  $a_n(u) = \nu_d(A_u(X, W_n))$  is the right statistic to study!

## Motivation: limiting variance in FCLT

By FCLT (Meschenmoser, Shashkin, 2011) and the continuous mapping theorem, it holds for some stationary weakly dependent associated random functions  $X$  with  $W_n = [0, n]^d$  that

$$\frac{\int_{\mathbb{R}} a_n(u) \mu(du) - n^d \int_{\mathbb{R}} \bar{F}_X(u) \mu(du)}{n^{d/2}} \xrightarrow{d} N(0, \sigma_{\mu, X}^2)$$

as  $n \rightarrow \infty$  for any finite measure  $\mu$  with  $\sigma_{\mu, X}^2$  as above.

So  $X$  is **SRD** if the asymptotic covariance  $\sigma_{\mu, X}^2$  in the CLT is finite for any finite measure  $\mu$  prescribing the choice of levels  $u$ .

## Motivation: American options

Let  $X = \{X_t, t \in \mathbb{Z}\}$  be the stock for which an American option at price  $u_0 > 0$ ,  $t \in [0, n]$ ,  $n \in \mathbb{N}$  is issued. The customer may buy a call at price  $u_0$  whenever  $X_t > u_0$  for some  $t \in [0, n]$ .

For  $\mu = \delta_{\{u_0\}}$  we get

$$\frac{\nu_1(\{t \in [0, n] : X_t > u_0\}) - n\bar{F}_X(u_0)}{\sqrt{n}} \xrightarrow{d} N(0, \sigma_{\delta_{\{u_0\}}, X}^2).$$

Then

- ▶  $X$  l.r.d. (i.e.,  $\sigma_{\delta_{\{u_0\}}, X}^2 = +\infty$ )  $\implies$  the amount of time within  $[0, n]$  at which the option may be exercised is not asymptotically normal for large time horizons  $n$ .
- ▶  $X$  s.r.d.  $\implies$  asymptotic normality of this time span for any price  $u_0$  provided that  $X$  satisfies some additional conditions.

## Motivation: Checking LRD

For a stationary centered Gaussian random function  $X$  with  $\text{Var } X_0 = 1$  and correlation function  $\rho(t)$  we have (Bulinski, S., Timmermann, 2012)

$$\text{cov}_X(t, u, v) = \frac{1}{2\pi} \int_0^{\rho(t)} \frac{1}{\sqrt{1-r^2}} \exp \left\{ -\frac{u^2 - 2ruv + v^2}{2(1-r^2)} \right\} dr.$$

## Motivation: statistical inference of LRD

The new definition is statistically feasible. Notice that for  $\mu = \delta_{\{u_0\}}$

$$\sigma_{\mu, X}^2 = \int_T |F_{X_0, X_t}(u_0, u_0) - F_X(u_0)F_X(u_0)| dt,$$

where the bivariate d.f.  $F_{X_0, X_t}(u, v) = P(X_0 \leq u, X_t \leq v)$  and marginal d.f.  $F_X(u) = P(X_0 \leq u)$  can be estimated from the data by their empirical counterparts.

## Motivation: LRD is margin-free

### Lemma (Kulik, S. 2019)

*A stationary real-valued random function  $X$  is SRD if*

$$\int_T \int_{[0,1]^2} |C_{0,t}(x, y) - xy| P_0(dx) P_0(dy) dt < +\infty$$

*for any probability measure  $P_0$  on  $[0, 1]$  where  $C_{0,t}$  is a copula of the bivariate distribution of  $(X_0, X_t)$ ,  $t \in T$ .  $X$  is LRD if there exists a probability measure  $P_0$  on  $[0, 1]$  such that the above integral is infinite.*

## Motivation: Checking LRD

Denote by  $P_\mu(\cdot) = \mu(\cdot)/\mu(\mathbb{R})$  the probability measure associated with the finite measure  $\mu$  on  $\mathbb{R}$ . If  $X \in \mathbf{PA}$  then applying Fubini–Tonelli theorem leads to

$$\sigma_{\mu, X}^2 = \mu^2(\mathbb{R}) \int_T \text{cov}(F_\mu(X_0), F_\mu(X_t)) dt,$$

where  $F_\mu(x) = P_\mu((-\infty, x))$  is the (left–side continuous) distribution function of probability measure  $P_\mu$ .

## Mixing

Let  $(\Omega, \mathcal{A}, P)$  be a probability space and  $(\mathcal{U}, \mathcal{V})$  be two sub- $\sigma$ -algebras of  $\mathcal{A}$ .  $\alpha$ -mixing coefficient:

$$\alpha(\mathcal{U}, \mathcal{V}) = \sup \{ |P(U \cap V) - P(U)P(V)| : U \in \mathcal{U}, V \in \mathcal{V} \}.$$

Let  $X = \{X_t, t \in T\}$  be a random function, and  $T$  be a normed space with distance  $d$ . Let  $X_C = \{X_t, t \in C\}$ ,  $C \subset T$ , and  $\mathcal{X}_C$  be the  $\sigma$ -algebra generated by  $X_C$ . If  $|C|$  is the cardinality of a finite set  $C$ , for any  $z \in \{\alpha, \beta, \phi, \psi, \rho\}$  put

$$z_X(k, u, v) = \sup \{ z(\mathcal{X}_A, \mathcal{X}_B) : d(A, B) \geq k, |A| \leq u, |B| \leq v \},$$

where  $u, v \in \mathbb{N}$  and  $d(A, B)$  is the distance between subsets  $A$  and  $B$ .



## SRD and mixing

### Theorem (Kulik, S. 2019)

Let  $X = \{X_t, t \in T\}$  be a stationary random function with  $z$ -mixing rate satisfying  $\int_T z_X(\|t\|, 1, 1) dt < +\infty$  where  $z \in \{\alpha, \beta, \phi, \psi, \rho\}$ . Then  $X$  is SRD with

$$\int_T \int_{\mathbb{R}^2} |\text{cov}_X(t, u, v)| \mu(du) \mu(dv) dt \leq 8 \int_T z_X(\|t\|, 1, 1) dt \cdot \mu^2(\mathbb{R})$$

for any finite measure  $\mu$ .

## Random volatility functions

Let the random function  $X = \{X_t, t \in T\}$  be given by

$$X_t = F(Y_t)Z_t$$

where  $Y = \{Y_t, t \in T\}$  and  $Z = \{Z_t, t \in T\}$  are independent stationary random functions,  $Z$  has property

$$\text{cov}_Z(t, u, v) \geq 0 \text{ or } \leq 0 \text{ for all } t \in T, u, v \in \mathbb{R},$$

$F : \mathbb{R} \rightarrow \mathbb{R}_\pm$  and  $P(F(Y_t) = 0) = 0$  for all  $t \in T$ .

$F(Y_t)$  is called a **random volatility** (being a deterministic function of a random (often LRD) function  $Y = \{Y_t, t \in T\}$ ) scaling a heavy tailed random function  $Z = \{Z_t, t \in T\}$ .

## Random volatility functions

### Theorem (Kulik, S. 2019)

Let the random volatility model  $X$  be given by  $X_t = AZ_t$ ,  $t \in T$ ,  $|T| = +\infty$  where  $A > 0$  a.s.,  $A$  and  $Z$  are independent and  $Z \in \mathbf{PA}$  is stationary. Then  $X$  is **LRD** if there exists  $u_0 \in \mathbb{R}$ :  $\bar{F}_Z(u_0/A) \neq \text{const}$  a.s.

## Random volatility functions

### Example

The above theorem evidently holds true if e.g.

- ▶  $Z_0 \sim \text{Exp}(\lambda)$ ,  $A \sim \text{Fréchet}(1)$  for any  $\lambda > 0$ .
- ▶  $X$  is a **subgaussian** random function where  $A = \sqrt{B}$ ,  $B \sim S_{\alpha/2} \left( \left( \cos \frac{\pi\alpha}{4} \right)^{2/\alpha}, 1, 0 \right)$ ,  $\alpha \in (0, 2)$ , and  $Z$  is a centered stationary Gaussian random function with covariance function  $C(t) \geq 0$  for all  $t \in T$  and a non-degenerate tail  $\bar{F}_Z$ .

## Random volatility functions

### Corollary

For the random function  $X = \{X_t, t \in T\}$  given by  $X_t = Y_t Z_t$ ,  $t \in T$ , assume that random functions  $Y = \{Y_t, t \in T\}$  and  $Z = \{Z_t, t \in T\}$  are stationary and independent. Assume that  $Z_0$  has a regularly varying tail, that is,  $P(Z_0 > x) \sim L(x)/x^\alpha$  as  $x \rightarrow +\infty$  for some  $\alpha > 0$  where the function  $L$  is slowly varying at  $+\infty$ . For  $Y_0 > 0$  a.s. assume that  $E Y_0^\delta < \infty$  and  $E(Y_0^\delta Y_t^\delta) < \infty$  for some  $\delta > \alpha$  and all  $t \in T$ . Let  $Y, Z \in \mathbf{PA}(\mathbf{NA})$ . Then  $X$  is **LRD** if  $Y^\alpha = \{Y_t^\alpha, t \in T\}$  is LRD.

## Random volatility functions

### Example

Assume that  $X_t = e^{Y_t^2/4} Z_t$ ,  $t \in \mathbb{Z}$ , where

- ▶  $Z_t$  is a sequence of i.i.d. random variables with finite moment of order  $2 + \delta$  for some  $\delta > 0$ ,
- ▶  $Y_t$  is a centered stationary Gaussian **PA** long memory sequence with unit variance and covariance function  $\rho$ ,
- ▶ sequences  $Z_t$  and  $Y_t$  are independent .

It holds  $EX_0^2 = +\infty$ . Choose  $\mu = \delta_{\{u_0\}}$  for some  $u_0 \in \mathbb{R}$ . Then

$$\sum_{t=1}^{\infty} \text{cov}_X(t, u_0, u_0) = \sum_{k=1}^{\infty} \frac{\langle \bar{F}_Z(u_0/G), H_k \rangle_{\varphi}^2}{k!} \sum_{t=1}^{\infty} \rho^k(t),$$

where  $G(x) = e^{x^2/4}$ .  $X$  is **LRD** if  $\sum_{t=1}^{\infty} \rho^2(t) = +\infty$ . In particular, if  $\rho(t) \sim |t|^{-\eta}$  as  $|t| \rightarrow \infty$ , then LRD occurs if  $\eta \in (0, 1/2]$ .

## LT for the volume of excursion sets

Let  $X$  be a real-valued random function on  $\mathbb{Z}^d$ ,  $d \geq 1$  and let  $W \subset \mathbb{Z}^d$  be a finite subset. Let

$$A_u(X, W) := \{t \in W : X(t) \geq u\}$$

be the **excursion set** of  $X$  in  $W$  over the level  $u \in \mathbb{R}$ .

**Asymptotic (non)Gaussian behavior of  $|A_u(X, W)|$  as  $W$  expands to  $\mathbb{Z}^d$ ?**

Prove a more general limit theorem for sums  $\sum_{t \in W} g(X_t)$  of functionals  $g$  of  $X$ !

## LT for the volume of excursion sets

Let  $X$  be a random volatility function of the form

$$X_t = G(Y_t)Z_t, \quad t \in \mathbb{Z}^d,$$

where

- ▶  $\{G(Y_t), t \in \mathbb{R}^d\}$  is a subordinated measurable Gaussian random function,
- ▶  $\{Z_t, t \in \mathbb{Z}^d\}$  is a white noise,
- ▶ the random functions  $Y$  and  $Z$  are independent.

Let  $W_n = [-n, n]^d$ , and  $g$  be a real valued function such that  $E[g(X_0)] = 0$ ,  $E[g^2(X_0)] > 0$ . Introduce the function

$$\xi(y) = E[g(G(y)Z_0)].$$

It follows that  $\xi(y) < \infty$  for  $\nu_1$ -a. e.  $y \in \mathbb{R}$ ,  $E[\xi(Y_0)] = 0$ .



## LT for the volume of excursion sets

Furthermore, set

$$m(y, Z_t) = g(G(y)Z_t) - \xi(y), \quad \chi(y) = E[m^2(y, Z_0)].$$

Assume that

- ▶  $\text{rank}(\xi) = q$ ,  $E[|g(X_0)|^2] < \infty$ ,  $E[\chi^3(Y_0)] < \infty$ .
- ▶  $Y$  is a homogeneous isotropic centered Gaussian random function with the covariance function  $\rho(t) = E[Y_0 Y_t] = |t|^{-\eta} L(|t|)$ ,  $\eta \in (0, d/q)$  and  $L$  is slowly varying at infinity,
- ▶  $Y$  has a spectral density  $f(\lambda)$  which is continuous for all  $\lambda \neq 0$  and decreasing in a neighborhood of 0.

## LT for the volume of excursion sets

## Theorem (Kulik, S. 2019)

1. If  $\xi(y) \equiv 0$  then

$$n^{-d/2} \sum_{t \in [-n, n]^d \cap \mathbb{Z}^d} g(X_t) \xrightarrow{d} \mathcal{N}(0, \sigma^2), \quad n \rightarrow +\infty,$$

where  $\sigma^2 = \mathbb{E}[g^2(X_0)]2^d > 0$ .

2. If  $\xi(y) \not\equiv 0$  then

$$n^{q\eta/2-d} L^{-q/2}(n) \sum_{t \in [-n, n]^d \cap \mathbb{Z}^d} g(X_t) \xrightarrow{d} R, \quad n \rightarrow +\infty,$$

where the random variable  $R$  is a  $q$ -Rosenblatt-type random variable.

## LT for the volume of excursion sets

$q$ -Rosenblatt-type random variable:

$$R = (\gamma(d, \eta))^{q/2} \int'_{\mathbb{R}^{dq}} \int_{[-1,1]^d} e^{i\langle \lambda_1 + \dots + \lambda_q, u \rangle} du \frac{B(d\lambda_1) \dots B(d\lambda_q)}{(|\lambda_1| \cdot \dots \cdot |\lambda_q|)^{(d-\eta)/2}},$$

$$\gamma(d, \eta) = \frac{\Gamma((d-\eta)/2)}{2^\eta \pi^{d/2} \Gamma(\eta/2)},$$

and  $\int'_{\mathbb{R}^{dq}}$  is the multiple Wiener–Ito integral with respect to a complex Gaussian white noise measure  $B$  (with structural measure being the spectral measure of  $Y$ ).

## LT for the volume of excursion sets

### Example

Assume that

$$g(y) = \mathbb{1}\{y > u\} - P(G(Y_0)Z_0 > u)$$

where  $G$  is nonnegative or nonpositive  $\nu_1$ -a.e. Then

$$\xi(y) = E[\mathbb{1}\{G(y)Z_0 > u\}] - P(G(Y_0)Z_0 > u).$$

- ▶ If  $u = 0$  then  $\xi(y) \equiv 0$ , so the Gaussian case applies.
- ▶ If  $u \neq 0$  then  $\xi(y) \not\equiv 0$ , so the non-Gaussian case applies.

Let  $uG(y) \geq 0$  for all  $y$ .

$q = 1$ :  $G : \mathbb{R} \rightarrow \mathbb{R}_\pm$  is monotone right-continuous non-constant fct. with  $\nu_1(\{x \in \mathbb{R} : G(x) = 0\}) = 0$ .

$q = 2$ :  $G(y) = G_1(|y|)$  with  $G_1$  as above.

## LT for the volume of excursion sets

### Example

Let the random volatility function  $X_t = G(|Y_t|)Z_t$ ,  $t \in \mathbb{Z}^d$  be s.t.

- ▶  $Y$  is a centered Gaussian random function with unit variance and corr. function  $\rho(t) \geq 0$  as above,  $\rho(t) \sim |t|^{-\eta}$  as  $|t| \rightarrow +\infty$
- ▶  $G(x) \geq 0$  is continuous as above with  $E |G(|Y_0|)|^{1+\theta} < \infty$  for some  $\theta \in (0, 1)$ .
- ▶  $\{Z_t\}$  is a heavy-tailed white noise,  $E Z_0^2 = +\infty$ .

## LT for the volume of excursion sets

For  $\tilde{G}(y) = G(|y|)$  and  $\mu = \delta_{\{u_0\}}$ ,  $u_0 > 0$  we have

$$\sum_{t \in \mathbb{Z}^d, t \neq 0} \text{cov}_X(t, u_0, u_0) = \sum_{k=1}^{\infty} \frac{\langle \bar{F}_Z(u_0/\tilde{G}), H_k \rangle_{\varphi}^2}{k!} \sum_{t \in \mathbb{Z}^d, t \neq 0} \rho^k(t),$$

- ▶ Since  $\text{rank}(\tilde{G}) = 2$ ,  $X$  is **LRD** if  $\sum_{t \in \mathbb{Z}^d, t \neq 0} \rho^2(t) = +\infty$ , that is, if  $\eta \in (0, d/2)$ .
- ▶ For niveau  $u \neq 0$ , the asymptotic behavior of  $|A_u(X, [-n, n]^d)|$  is of 2-Rosenblatt-type ( $\text{rank}(\xi) = q = 2$ ) if  $\eta \in (0, d/2)$ .

## LT for the volume of excursion sets

### Summary:

The correct statistics associated with the new definition of l.r.d. is the volume of excursion sets!!!!

## Linear $\alpha$ -stable time series

- ▶ Let  $\{Z_t, t \in \mathbb{Z}\}$  be a sequence of i.i.d.  $S_{\alpha}S$  random variables with characteristic function

$$\psi_Z(\mathbf{s}) = \exp\{-\tau^\alpha |\mathbf{s}|^\alpha\}$$

for  $\tau > 0$ ,  $\alpha \in (1, 2)$ ,  $\mathbf{s} \in \mathbb{R}$ .

- ▶ Let  $\{a_j, j \in \mathbb{Z}\}$  be a nonnegative number sequence s. t.

$$\sum_{j=-\infty}^{+\infty} a_j < \infty$$

- ▶ **Linear  $S_{\alpha}S$  time series:**

$$Y(t) = \sum_{j=-\infty}^{+\infty} a_j Z_{t-j}, \quad t \in \mathbb{Z}.$$



## SRD/LRD for linear $\alpha$ -stable time series

Let  $\{Y(t) = \sum_{j=-\infty}^{+\infty} a_j Z_{t-j}, \quad t \in \mathbb{Z}\}$  be as above.

**Theorem (Makogin, Oesting, Rapp, S. (2019))**

- ▶  $Y$  is **SRD** if  $\sum_{j=-\infty}^{\infty} a_j^{\alpha/2} < \infty$ .
- ▶  $Y$  is **LRD** if  $\sum_{j=-\infty}^{\infty} \sum_{t=-\infty}^{\infty} (a_j^{\alpha} \wedge a_t^{\alpha}) = \infty$ .
- ▶ If  $a_j$  is monotonically decreasing on  $\mathbb{Z}_+$  and  $a_j = a_{-j}$  for all  $j \in \mathbb{Z}$  then  $Y$  is **LRD** whenever  $\sum_{t=0}^{\infty} t a_t^{\alpha} = \infty$ .

## Max-stable stationary processes

- ▶ A stochastic process  $X = \{X(t), t \in T\}$  is called **max-stable** if, for all  $n \in \mathbb{N}$ , there exist functions  $a_n : T \rightarrow (0, \infty)$  and  $b_n : T \rightarrow \mathbb{R}$  such that

$$\left\{ \max_{i=1, \dots, n} \frac{X_i(t) - b_n(t)}{a_n(t)}, t \in T \right\} \stackrel{d}{=} \{X(t), t \in T\},$$

where the processes  $X_i, i \in \mathbb{N}$ , are independent copies of  $X$ .

- ▶ **Marginal distributions** of a max-stable process: degenerate, Fréchet, Gumbel or Weibull law.
- ▶  **$\alpha$ -Fréchet marginal distribution**:  $P(X(t) \leq x) = \exp(-x^{-\alpha})$  for all  $x > 0$  and some  $\alpha > 0$  and all  $t \in T$ . Here, **covariances do not exist** if  $\alpha \leq 2$ .

## Max-stable stationary processes

- ▶ **Pairwise extremal coefficient:**  $\{\theta_t, t \in T\}$  defined via

$$P(X(0) \leq x, X(t) \leq x) = P(X(0) \leq x)^{\theta_t} \quad \text{for all } x > 0,$$

- ▶ It holds  $\theta_t = 2 - \lim_{x \rightarrow \infty} P(X(t) > x \mid X(0) > x)$ .
- ▶  $\theta_t \in [1, 2]$ , where
  - ▶  $\theta_t = 2 \implies X(0)$  and  $X(t)$  asymptotically independent,
  - ▶  $\theta_t = 1 \implies X(0)$  and  $X(t)$  asymptotically fully dependent.

## SRD/LRD for max-stable stationary processes

Theorem (Makogin, Oesting, Rapp, S. (2019))

Let  $X = \{X(t), t \in T\}$  be a stationary max-stable process with  $\alpha$ -Fréchet marginal distribution and pairwise extremal coefficient  $\{\theta_t, t \in T\}$ .  $X$  is **LRD** iff

$$\int_T (2 - \theta_t) dt = \infty.$$

## Outlook

- ▶ Checking the new LRD definition for other classes of processes with infinite variance, e.g., for infinitely divisible moving averages
- ▶ Connection of LRD with LT for the volume of excursions of other stationary random functions

## Literature: Historical and general expositions

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## Appendix: Subordinated Gaussian random function

Let  $Y = \{Y_t, t \in T\}$  be a stationary centered Gaussian real-valued random function with  $\text{Var}(Y_t) = 1$  and  $\rho(t) = \text{Cov}(Y_0, Y_t)$ ,  $t \in T$ . The **subordinated Gaussian random function**  $X$  is defined by

$$X_t = g(Y_t), t \in T,$$

where  $g : \mathbb{R} \rightarrow \text{Im}(g) \subseteq \mathbb{R}$  is a measurable function.



## Expansions in Hermite polynomials

Let  $\varphi(x)$  be the density and  $\Phi(x)$  the c.d.f. of the standard normal law. **Hermite polynomials**  $H_n$

- ▶ are defined by  $H_n(x) = (-1)^n \frac{\varphi^{(n)}(x)}{\varphi(x)}$ ,  $n = 0, 1, 2, \dots$
- ▶ are polynomials of degree  $n$ :  $H_0(x) = 1$ ,

$$H_1(x) = x, \quad H_2(x) = x^2 - 1, \quad H_3(x) = x^3 - 3x, \dots$$

- ▶ form an orthogonal basis of the Hilbert space of square integrable with  $e^{-\frac{x^2}{2}}$  functions with inner product

$$\langle f, g \rangle_\varphi = \int_{-\infty}^{+\infty} f(x)g(x)\varphi(x) dx.$$

Hence, any function from this space has a series expansion w.r.t. Hermite polynomials.

## Appendix: Expansions in Hermite polynomials

### Lemma (Rozanov (1967))

Let  $Z_1, Z_2$  be standard normal random variables with  $\rho = \text{cov}(Z_1, Z_2)$ , and let  $G$  be a function satisfying  $E[G(Z_1)] = 0$  and  $E[G^2(Z_1)] < +\infty$ . Then

$$\text{Cov}(G(Z_1), G(Z_2)) = \sum_{k=1}^{\infty} \frac{\langle G, H_k \rangle_{\varphi}}{k!} \rho^k.$$

Assume  $Y = \{Y_t, t \in T\}$  to be a stationary centered Gaussian real-valued random function with  $\text{Var}(Y_t) = 1$  and  $\rho(t) = \text{Cov}(Y_0, Y_t)$ . **Classical definition of LRD of  $X = g(Y)$**  with  $C(t) = \text{Cov}(X_0, X_t) \geq 0, t \in T$  yields

$$\int_T |C(t)| dt = \int_T \sum_{k=1}^{\infty} \frac{\langle G, H_k \rangle_{\varphi}}{k!} \rho(t)^k dt = \sum_{k=1}^{\infty} \frac{\langle G, H_k \rangle_{\varphi}}{k!} \int_T \rho(t)^k dt = +\infty.$$

## Appendix: Subordinated Gaussian random functions

Let  $T \subseteq \mathbb{R}^d$ , and  $\nu_d$  be the Lebesgue measure on  $\mathbb{R}^d$ .

### Theorem (Kulik, S., 2019)

Let  $X$  be a subordinated Gaussian random function defined by  $X_t = g(Y_t)$ ,  $t \in T$ , where  $g$  is a right-continuous strictly monotone (increasing or decreasing) function. Assume  $\nu_d(\{t \in T : \rho(t) = 1\}) = 0$ . Let

$$b_k(\mu) = \left( \int_{\text{Im}(g)} H_k(g^{-}(u)) \varphi(g^{-}(u)) \mu(du) \right)^2$$

where  $g^{-}$  is the generalized inverse of  $g$  if  $g$  is increasing or of  $-g$  if  $g$  is decreasing. Then  $X$  is **SRD** if for any finite measure  $\mu$

$$\sum_{k=1}^{\infty} \frac{b_{k-1}(\mu)}{k!} \int_T |\rho(t)| \rho(t)^{k-1} dt < +\infty.$$

## Appendix: Subordinated Gaussian random functions, Remarks

- ▶ If  $X_t = g(|Y_t|)$ ,  $t \in T$ , then the above **SRD condition** modifies to

$$\sum_{k=1}^{\infty} \frac{b_{2k-1}(\mu)}{(2k)!} \int_T \rho(t)^{2k} dt < +\infty. \quad (2)$$

- ▶ **LRD conditions can be formulated:** e.g.,  $X$  is LRD if  $b_k(\delta_{u_0}) < +\infty$  for some  $u_0 \in \mathbb{R}$  and all  $k$ , and the above series diverges to  $+\infty$ .

## Appendix: Subordinated Gaussian random functions, Example

- ▶ Let  $g(x) = e^{x^2/(2\alpha)}$ ,  $T = \mathbb{R}^d$ ,  $\alpha > 0$ .  
For  $\alpha \in (1, 2]$ ,  $E X_0 < \infty$ , but  $E X_0^2 = +\infty$ .
- ▶ One can show that  $\frac{b_{2k-1}(\mu)}{(2k)!} = O\left(\frac{1}{\sqrt{k}}\right)$ ,  $k \rightarrow +\infty$ .
- ▶ For  $\rho(t) \sim |t|^{-\eta}$  as  $|t| \rightarrow +\infty$ ,  $\eta > 0$ ,  $X = e^{Y^2/(2\alpha)}$  is
  - ▶ **LRD** if  $\eta \in (0, d/2]$ , since then  $\int_{\mathbb{R}^d} \rho^2(t) dt = +\infty$ ,
  - ▶ **SRD** if  $\eta > d/2$ , since

$$\int_{\mathbb{R}^d} \rho^{2k}(t) dt = O(k^{-1}) \quad \text{as } k \rightarrow +\infty,$$

and the series (2) behaves as

$$\sum_{k=1}^{\infty} \frac{1}{k^{3/2}} < +\infty.$$

- ▶ Hence, for  $\eta \in (d/2, d)$  **Y is LRD but  $X = e^{Y^2/(2\alpha)}$  is SRD!**